6 APPENDIX 1

## SUPPLMENTARY MATERIAL FOR

## COONEY ET AL. - MULTI-MODAL SIGNAL EVOLUTION IN BIRDS: RE-EXAMINING A STANDARD PROXY FOR SEXUAL SELECTION

Supplementary methods

Supplementary figures (S1-S10) and tables (S1-S12)

## APPENDIX 1

## SUPPLMENTARY METHODS

## Compiling species pair datasets

Following Seddon et al. (2013) and Cooney et al. (2017), we assembled a dataset of passerine species pairs for which we could collect detailed data on song traits. We compiled a list of species pairs from published phylogenetic trees of passerine families or genera generated using genetic data. We only included trees in which (i) > 70\% of taxa had been sampled, and (ii) node support was high, with either posterior probability > $95 \%$ or ML bootstrap support > $70 \%$. When several phylogenies were presented in a paper, we only selected species pairs resolved in all trees. When nodal support varied with the method of phylogenetic reconstruction, ML bootstrap values took precedence. We assumed that consensus trees and trees based on concatenated molecular datasets provided the most reliable source of phylogenetic information and thus, whenever possible, we assessed nodal support based on the values given in these trees.

## Principal component analysis of song measurements

Raw song traits were correlated across species (Table S1). Such correlations can lead to overinflated divergence estimates if each trait axis is treated as independent. Therefore, to avoid this issue and to reduce the dimensionality of our dataset, following previous studies (Tobias et al. 2014a; Tobias et al. 2014b) we conducted a principal components (PC) analysis on the covariance matrix of individual (log-transformed) song measurements. The first three PCs from this analysis accounted for over $83 \%$ of the variance in the original acoustic dataset, with each PC capturing a distinct component of overall signal structure (Table S1). Specifically, PC1 (41\%) primarily loaded with variables related to song pitch, PC2 (24\%) with variables related to song length, and PC3 (18\%) with song pace. We chose not to retain PC4 (14\%) in our analysis because of its comparatively small eigenvalue (1.006) (Kaiser 1960) and because the loadings indicated that it was primarily related to variation in minimum song frequency (loading $=0.66$ ), which is more difficult to quantify from song spectra and therefore potentially subject to greater levels of error. By focusing exclusively on variation in the first three PCs, which are interpretable in terms of song pitch (PC1), length (PC2) and pace (PC3), respectively, our analyses are directly comparable with previous studies (e.g. Tobias et al. 2014b) using the same combination of traits (PCs) to define a threedimensional acoustic space (Fig. S2).

## Alternative dataset: accounting for intraspecific variation

Intraspecific trait variation can generate overinflated estimates of interspecific disparity, especially when sample sizes are low relative to levels of intraspecific variation and/or measurement error
(Silvestro et al. 2015). To test whether our results were sensitive to these issues, we repeated our main analyses using an alternative version of our dataset in which within-species song disparity estimates were corrected for observed levels of intraspecific variation using an approach based on ANOVA (Weir and Wheatcroft 2011; Weir et al. 2012).

## Sexual dichromatism

Although it would be preferable to estimate dichromatism using more objective measurements of bird plumage colouration, such as those generated by a spectrophotometer, these types of measurements are currently unavailable for most species. However, the high correlation between spectrophotometric and human estimates of dichromatism indicate that human vision can provide a reliable estimate of avian colouration in general (Bergeron and Fuller 2018) and sexual dichromatism in particular (Armenta et al. 2008; Seddon et al. 2010). To test this assumption in our dataset, we used spectrophotometric measurements of plumage colouration from museum specimens described in (Cooney et al. 2017) to calculate objective estimates of dichromatism for species and pairs common to both datasets, following methods outlined previously (Seddon et al. 2013; Cooney et al. 2017). This resulted in dichromatism estimates for 264 species ( 132 pairs), which we compared to our scores based on handbook illustrations using Spearman's rank correlation. We found that estimates of dichromatism based on spec measurements and human vision were highly positively correlated (Fig. S9) across both species (Spearman's $r=0.76, P<$ 0.001 ) and species pairs (Spearman's $r=0.79, P<0.001$ ). Thus, we conclude that scores of dichromatism based on published illustrations are valid for our dataset, in line with previous studies (e.g. Dale et al. 2015).

## Additional predictors of song divergence

To explore the role of other factors known to influence song divergence in birds, we collected data for a suite of additional explanatory variables. Details are given below.

Pair age. We generated estimates of pair age (i.e. time from the present to the most recent common ancestor) by calculating the mean height of the node connecting the pair in the (pruned) Jetz et al. (2012) phylogenies (see above).

Body mass. Estimates of body mass were extracted from Wilman et al. (2014) and we used the mean of species (In-transformed) values to generate pair-level averages.

Generation length. Data on species' generation length, forest dependency and migratory status were compiled from BirdLife International (http://www.datazone.birdlife.org; accessed 18-102017), and we used the mean of species' generation length values to provide pair-level averages.

Forest dependency. In the BirdLife dataset species are assigned to one of four categories, depending on whether they "do not normally occur in forests", or exhibit "low", "medium" or "high" levels of forest dependency. To convert this into a quantitative variable capturing broad differences
in habitat usage, we converted this classification system into a 4-point scoring scheme, giving each species a score of $0,1,2$ or 3 based on whether they were recorded as having no, low, medium or high forest dependency, respectively. To capture the average level of forest dependency within a pair, we took the mean score for each pair.

Migratory status. Similarly, in the BirdLife dataset species are categorised as "not a migrant", "nomadic", "altitudinal migrant" or "full migrant", so again we converted this classification system into a 2-point scoring scheme, giving each species a score of 0 (not a migrant) or 1 (nomadic, altitudinal migrant or full migrant) and then took the mean score for each pair.

Latitude. All variables relating to species' geographical distributions are based on (breeding) range maps provided by BirdLife International and NatureServe (2016, version 9; http://www.datazone.birdlife.org). To estimate the latitudinal midpoint of each pair, we first calculated each species' range centroids using the R package PBSmapping (Schnute 2015) and then took the mean of species (absolute) latitudinal values (Weir and Schluter 2007; Weir and Wheatcroft 2011; Weir et al. 2012).

Island-dwelling. To characterise species as island-dwelling, we used a high-resolution database of global geography (GSHHG version 2.3.6; http://www.soest.hawaii.edu/pwessel/gshhg/) to identify all landmasses $>1 \mathrm{~km}^{2}$ in size surrounded by ocean and smaller than Greenland (following Weigelt et al. 2013). We then overlaid species range maps on top of this dataset and calculated the proportion of species' ranges occurring on islands. We consider species with $>80 \%$ of their range occurring on islands to be primarily island dwelling and turned this into a pair-level variable by counting whether neither (0), one (1) or both (2) of the species in a pair were characterised as island dwelling.

Within-pair range overlap. To calculate range overlap between species within a pair, we took the standard approach of calculating the area of overlap between species, divided by the area of the smaller species' range (Pigot and Tobias 2013; Tobias et al. 2014a; Cooney et al. 2017).

Confamilial sympatry. To quantify the extent of interactions with related species, we used the taxonomy of Jetz et al. (2012) and the range maps to tally up the number of same-family (confamilial) species that are sympatric with the focal species (Morinay et al. 2013), where sympatry is defined as $>20 \%$ range overlap between species-a threshold commonly used to define 'substantial' sympatry (Tobias et al. 2014a). We then divided this number by the total number of species in the family and took the mean of these values to generate a pair-level average.

Climatic disparity. To quantify divergence in species' climatic niches, we extracted rangewide climatic information for each species from WorldClim (http://www.worldclim.org/) using a global resolution of 10 arc minutes $\left(\sim 100 \mathrm{~km}^{2}\right)$. Following Lawson and Weir (2014), we focused on 48 climatic variables, 36 related to temperature (maximum, minimum and mean for each month of the year) and 12 to precipitation (mean for each month of the year). We then conducted a PC analysis on these (sorted) values across all species (Lawson and Weir 2014), retaining the first four PC axes
that together accounted for $96 \%$ of the variation. Euclidean distances between the species' midpoint (median) values in each of the four axes were used as an overall measure of climatic divergence within a pair.

Body mass disparity. To calculate within-pair disparity in body mass, we took the absolute difference between (In-transformed) body mass values for species within a pair.

Beak disparity. Variables relating to species' beak traits are based on a dataset of biometric measurements collected from museum specimens. For each species, we measured linear measurements of beak morphology (length, width and depth, measures at the nares) for multiple individuals per species (mean = 6.5) per sex. To quantify divergence in beak morphology, we followed the methods outlined in Pigot and Tobias (2013) and calculated beak divergence as the Euclidean distance between species mean (In-transformed) values for all three axes.

Song learning. Finally, we identified whether species pairs belonged to the oscine (Passeri) or suboscine (Tyranni) passerine suborders, which do and do not exhibit song learning, respectively.

## Collinearity among predictor variables

Collinearity among predictor variables is a potential problem in multiple regression modelling (Graham 2003). However, methods based on the type of information theoretic-based model averaging approach we employ here are generally robust to collinearity (Graham 2003; Freckleton 2011). Nonetheless, we checked for collinearity among predictors by calculating pairwise Pearson correlation coefficients ( $r^{2}$ ) and variance inflation factors (VIF) for all variables. Only one pair of predictors (latitude and migration) showed evidence of (strong) collinearity ( $r^{2}=0.75$; Table S12), and calculating VIFs for our full dataset revealed no evidence of severe (VIF > 10) or even moderate (VIF > 4) multicollinearity (range $=1.16-3.14$; mean $=1.61$ ), indicating that multicollinearity is unlikely to affect our results.

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Figure S1. A representative phylogeny of passerines from Jetz et al. (2012) showing the location of species sampled in this study (green points).


Pitch (PC1; 41\%)

Figure S2. Plot showing the distribution of individual song recordings ( $n=2476$ ) in acoustic 'trait space', a defined by the first three principal component (PC) axes of variation, corresponding to variation among song in pitch (PC1), length (PC2) and pace (PC3). Coloured areas represent kernel density estimates of $50 \%$ (red), $75 \%$ (orange) and $100 \%$ (yellow) of the data (for illustrative purposes only).


Figure S3. Histograms showing the distribution of within-pair song disparity estimates for species pairs $(\mathrm{n}=259)$.


Figure S4. Plots showing the relationship between overall within-pair song disparity (PC1-3) and each of the additional predictors included in our analysis. For all continuous predictors, regressions lines (with prediction intervals, shaded) indicate the best-fitting relationship between the two variables.


Figure S5. Plots showing the relationship between within-pair disparity in pitch (PC1) and each of the additional predictors included in our analysis. For all continuous predictors, regressions lines (with prediction intervals, shaded) indicate the best-fitting relationship between the two variables.


Figure S6. Plots showing the relationship between within-pair disparity in length (PC2) and each of the additional predictors included in our analysis. For all continuous predictors, regressions lines (with prediction intervals, shaded) indicate the best-fitting relationship between the two variables.


Figure S7. Plots showing the relationship between within-pair disparity in pace (PC3) and each of the additional predictors included in our analysis. For all continuous predictors, regressions lines (with prediction intervals, shaded) indicate the best-fitting relationship between the two variables.


Figure S8. Model averaged coefficient estimates from multipredictor GLMs predicting variation in within-pair disparity in song pitch (PC1), length (PC2) and pace (PC3) among passerine species pairs $(\mathrm{n}=246)$. Within each dataset, points indicate the standardised effect sizes for each of the (scaled) predictor variables and lines indicate $95 \%$ confidence intervals $(\mathrm{Cl})$. Sizes of points represent the relative importance (RI) of each of the predictor variables, where a value of $\mathrm{RI}=0$ indicates low importance and a value of $\mathrm{RI}=1$ indicates high importance. Predictors included in the $\mathrm{AIC}_{\mathrm{c}}$ top model for each dataset are coloured (blue = negative effect; red $=$ positive effect), with significant $(P<0.05)$ model-averaged coefficients shown in darker colours.


Figure S9. Scatterplots showing the relationship between dichromatism estimates based on spectrophotometric measurements of colouration from museum specimens and handbook illustrations for species $(n=264)$ and species pairs $(n=132)$ of birds.


Figure S10. A phylogenetic tree of passerine species pairs ( $\mathrm{n}=259$ ), showing within-pair disparity in overall song (PC1-3; black), pitch (PC2; red), length (PC3; blue) and pace (PC3; green).

|  | Pitch (PC1) | Length (PC2) | Pace (PC3) |
| :--- | :---: | :---: | :---: |
| Maximum frequency | $\mathbf{- 0 . 5 7}$ | 0.06 | 0.11 |
| Minimum frequency | -0.27 | -0.26 | 0.41 |
| Peak | $\mathbf{- 0 . 5 4}$ | -0.08 | 0.22 |
| Bandwidth | $\mathbf{- 0 . 4 8}$ | 0.16 | -0.07 |
| Duration | 0.12 | $\mathbf{0 . 6 4}$ | 0.45 |
| Note number | -0.09 | $\mathbf{0 . 7 0}$ | -0.16 |
| Pace | -0.25 | 0.07 | $\mathbf{- 0 . 7 3}$ |
| Eigenvalue | 2.85 | 1.72 | 1.25 |
| Cumulative variance explained (\%) | 41 | 65 | 83 |

Table S1. Variable loadings and variance explained by the first three principal components (PCs) of an analysis of measurements of individual song recordings ( $n=2476$ ) for the species included in our analyses ( $\mathrm{n}=518$ ). Standardised loadings of the main contributors to each component are highlighted in bold.

|  | Full dataset |  |  |  | Alternative dataset |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | $\lambda(95 \% \mathrm{Cl})$ | $P(\lambda=0)$ | $P(\lambda=1)$ |  | $\lambda(95 \% \mathrm{Cl})$ | $P(\lambda=0)$ | $P(\lambda=1)$ |
| Total disparity (PC1-3) | $0.00(\mathrm{NA}, 0.39)$ | 1.000 | $<0.001$ |  | $0.00(\mathrm{NA}, 0.29)$ | 1.000 | $<0.001$ |
| Pitch disparity (PC1) | $0.00(\mathrm{NA}, 0.09)$ | 1.000 | $<0.001$ |  | $0.00(\mathrm{NA}, 0.08)$ | 1.000 | $<0.001$ |
| Length disparity (PC2) | $0.00(\mathrm{NA}, 0.50)$ | 1.000 | $<0.001$ |  | $0.00(\mathrm{NA}, 0.14)$ | 1.000 | $<0.001$ |
| Pace disparity (PC3) | $0.00(\mathrm{NA}, 0.18)$ | 1.000 | $<0.001$ |  | $0.01(\mathrm{NA}, 0.30)$ | 0.719 | $<0.001$ |

Table S2. Maximum likelihood estimates (and 95\% confidence intervals; CI ) of phylogenetic signal (Pagel's lambda; $\lambda$ ) for each measure of within-pair song disparity across passerine species pairs ( $\mathrm{n}=259$ ).

| Response variable | Term | Full dataset |  |  | Alternative dataset |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Estimate (SE) | $P$ | $R^{2}$ | Estimate (SE) | $P$ | $R^{2}$ |
| Total disparity (PC1-3) | Dichromatism | -0.25 (0.08) | 0.002 | 0.04 | -0.16 (0.06) | 0.001 | 0.03 |
| Pitch disparity (PC1) | Dichromatism | -0.27 (0.11) | 0.015 | 0.03 | -0.13(0.05) | 0.015 | 0.04 |
| Length disparity (PC2) | Dichromatism | -0.33 (0.12) | 0.005 | 0.04 | -0.14 (0.06) | 0.019 | 0.03 |
| Pace disparity (PC3) | Dichromatism | -0.21 (0.12) | 0.074 | 0.02 | -0.08 (0.05) | 0.137 | 0.01 |
| Total disparity (PC1-3) | Dichromatism | -0.26 (0.08) | 0.001 | 0.06 | -0.17 (0.06) | 0.005 | 0.06 |
|  | Forest dependency | 0.17 (0.08) | 0.039 |  | 0.15 (0.06) | 0.013 |  |
|  | Interaction | -0.08 (0.15) | 0.585 |  | -0.05 (0.11) | 0.627 |  |
| Pitch disparity (PC1) | Dichromatism | -0.25 (0.10) | 0.009 | 0.07 | -0.14 (0.05) | 0.009 | 0.07 |
|  | Forest dependency | 0.30 (0.10) | 0.006 |  | 0.13 (0.05) | 0.011 |  |
|  | Interaction | -0.10 (0.21) | 0.633 |  | -0.02 (0.10) | 0.803 |  |
| Length disparity (PC2) | Dichromatism | -0.33 (0.12) | 0.005 | 0.04 | -0.14 (0.06) | 0.019 | 0.04 |
|  | Forest dependency | 0.02 (0.12) | 0.875 |  | 0.06 (0.06) | 0.311 |  |
|  | Interaction | -0.10 (0.22) | 0.641 |  | -0.06 (0.11) | 0.623 |  |
| Pace disparity (PC3) | Dichromatism | -0.22 (0.12) | 0.065 | 0.02 | -0.09 (0.05) | 0.116 | 0.02 |
|  | Forest dependency | 0.16 (0.12) | 0.181 |  | 0.06 (0.05) | 0.261 |  |
|  | Interaction | -0.06 (0.22) | 0.796 |  | 0.02 (0.10) | 0.820 |  |

Table S3. Standardised coefficient estimates from GLMs predicting variation in within-pair disparity in song traits among passerine species pairs ( $\mathrm{n}=259$ ). SE, standard error. $R^{2}$, pseudo- $R^{2}$.

Full dataset

|  | N pairs | Estimate (SE) | $P$ | $R^{2}$ |
| :--- | :---: | :---: | :---: | :---: |
| Plumage dichromatism | 259 | $-0.25(0.08)$ | 0.002 | 0.04 |
| Pair age | 259 | $0.34(0.08)$ | $<0.001$ | 0.09 |
| Body mass | 259 | $0.15(0.08)$ | 0.081 | 0.02 |
| Generation length | 251 | $0.08(0.08)$ | 0.334 | $<0.01$ |
| Forest dependency | 259 | $0.15(0.08)$ | 0.063 | 0.02 |
| Migratory status | 259 | $-0.06(0.08)$ | 0.467 | $<0.01$ |
| Latitude | 257 | $0.00(0.08)$ | 0.987 | $<0.01$ |
| Island dwelling | 259 | $0.12(0.08)$ | 0.150 | 0.01 |
| Within-pair range overlap | 257 | $0.05(0.08)$ | 0.565 | $<0.01$ |
| Confamilial sympatry | 257 | $-0.08(0.08)$ | 0.330 | 0.01 |
| Body mass disparity | 259 | $0.16(0.08)$ | 0.053 | 0.02 |
| Bill disparity | 257 | $0.25(0.08)$ | 0.003 | 0.04 |
| Climatic disparity | 254 | $0.10(0.08)$ | 0.252 | 0.01 |
| Oscine/suboscine | 259 | $0.01(0.08)$ | 0.919 | $<0.01$ |

Table S4. Standardised coefficient estimates from single predictor GLMs predicting variation in within-pair song disparity (PC1-3) among passerine species pairs ( $n=259$ ). SE, standard error. $R^{2}$, pseudo- $R^{2}$.

Full dataset

|  | N pairs | Pitch disparity (PC1) |  |  | Length disparity (PC2) |  |  | Pace disparity (PC3) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Estimate (SE) | $P$ | $R^{2}$ | Estimate (SE) | $P$ | $R^{2}$ | Estimate (SE) | $P$ | $R^{2}$ |
| Plumage dichromatism | 259 | -0.27 (011) | 0.015 | 0.03 | -0.33 (0.12) | 0.001 | 0.04 | -0.21 (0.12) | 0.074 | 0.02 |
| Pair age | 259 | 0.38 (0.12) | 0.001 | 0.07 | 0.37 (0.12) | 0.002 | 0.06 | 0.31 (0.12) | 0.010 | 0.04 |
| Body mass | 259 | 0.25 (0.11) | 0.028 | 0.03 | 0.20 (0.12) | 0.088 | 0.02 | -0.05 (0.12) | 0.677 | $<0.01$ |
| Generation length | 251 | 0.15 (0.12) | 0.203 | 0.01 | 0.11 (0.12) | 0.353 | 0.01 | 0.04 (0.12) | 0.741 | <0.01 |
| Forest dependency | 259 | 0.29 (0.11) | 0.008 | 0.04 | 0.01 (0.12) | 0.941 | <0.01 | 0.14 (0.12) | 0.229 | 0.01 |
| Migratory status | 259 | -0.21 (0.11) | 0.060 | 0.02 | 0.03 (0.12) | 0.783 | <0.01 | -0.02 (0.12) | 0.868 | <0.01 |
| Latitude | 257 | -0.15 (0.11) | 0.180 | 0.01 | 0.12 (0.12) | 0.299 | 0.01 | 0.05 (0.12) | 0.701 | <0.01 |
| Island dwelling | 259 | 0.19 (0.11) | 0.096 | 0.02 | 0.14 (0.12) | 0.234 | 0.01 | 0.08 (0.12) | 0.518 | $<0.01$ |
| Within-pair range overlap | 257 | 0.06 (0.11) | 0.621 | $<0.01$ | -0.01 (0.12) | 0.937 | <0.01 | 0.09 (0.12) | 0.448 | <0.01 |
| Confamilial sympatry | 257 | -0.12 (0.11) | 0.281 | 0.01 | -0.16 (0.12) | 0.182 | 0.01 | 0.01 (0.12) | 0.954 | <0.01 |
| Body mass disparity | 259 | 0.36 (0.11) | 0.002 | 0.06 | 0.11 (0.12) | 0.371 | <0.01 | 0.09 (0.12) | 0.441 | <0.01 |
| Bill disparity | 257 | 0.23 (0.11) | 0.047 | 0.02 | 0.26 (0.12) | 0.030 | 0.03 | 0.28 (0.12) | 0.018 | 0.03 |
| Climatic disparity | 254 | 0.04 (0.12) | 0.738 | <0.01 | 0.20 (0.12) | 0.086 | 0.02 | 0.12 (0.12) | 0.315 | 0.01 |
| Oscine/suboscine | 259 | -0.01 (0.11) | 0.919 | $<0.01$ | -0.14 (0.12) | 0.234 | 0.01 | 0.15 (0.12) | 0.215 | 0.01 |

Table S5. Standardised coefficient estimates from single predictor GLMs predicting variation in within-pair disparity in three independent axes of song variation among passerine species pairs $(\mathrm{n}=259)$. SE, standard error. $R^{2}$, pseudo- $R^{2}$.

|  | Model averaged |  |  |  | AIC $_{c}$ top model |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate (SE) |  | $P$ | RI |  | Estimate (SE) |
| Plumage dichromatism | $-0.25(0.09)$ | 0.004 | $0.98^{*}$ |  | $-0.25(0.08)$ | 0.003 |
| Pair age | $0.28(0.09)$ | 0.002 | $0.99^{*}$ |  | $0.26(0.09)$ | 0.002 |
| Body mass | $0.17(0.09)$ | 0.070 | $0.74^{*}$ |  | $0.16(0.09)$ | 0.057 |
| Generation length | $0.04(0.11)$ | 0.731 | 0.29 |  | - | - |
| Forest dependency | $0.15(0.09)$ | 0.106 | $0.61^{*}$ |  | $0.17(0.09)$ | 0.064 |
| Migratory status | $-0.01(0.11)$ | 0.959 | 0.26 |  | - | - |
| Latitude | $0.02(0.13)$ | 0.868 | 0.28 |  | - | - |
| Island dwelling | $0.05(0.09)$ | 0.576 | 0.29 |  | - | - |
| Within-pair range overlap | $0.06(0.09)$ | 0.492 | 0.31 |  | - | - |
| Confamilial sympatry | $-0.16(0.10)$ | 0.093 | $0.65^{*}$ |  | $-0.18(0.09)$ | 0.045 |
| Body mass disparity | $-0.02(0.10)$ | 0.854 | 0.27 |  | - | - |
| Bill disparity | $0.15(0.09)$ | 0.104 | $0.62^{*}$ |  | $0.14(0.09)$ | 0.098 |
| Climatic disparity | $0.03(0.09)$ | 0.737 | 0.27 |  | - | - |
| Oscine/suboscine | $0.00(0.09)$ | 0.968 | 0.26 |  | - | - |

Table S6. Standardised coefficient estimates from model averaged and $\mathrm{AIC}_{\mathrm{c}}$ top model multipredictor GLMs predicting variation in within-pair song disparity (PC1-3) among passerine species pairs ( $\mathrm{n}=246$ ). SE, standard error; RI, relative importance. Pseudo- $R^{2}$ for $\mathrm{AlC}_{\mathrm{c}}$ top model $=0.17$.

|  | Pitch disparity (PC1) |  |  | Length disparity (PC2) |  |  | Pace disparity (PC3) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate (SE) | $P$ | RI | Estimate (SE) | $P$ | Rl | Estimate (SE) | $P$ | RI |
| Plumage dichromatism | -0.33 (0.11) | 0.004 | 0.97* | -0.32 (0.12) | 0.010 | 0.91* | -0.14 (0.12) | 0.252 | 0.39 |
| Pair age | 0.33 (0.12) | 0.005 | 0.96* | 0.32 (0.12) | 0.009 | 0.92* | 0.20 (0.13) | 0.126 | 0.56* |
| Body mass | 0.32 (0.13) | 0.010 | 0.93* | 0.24 (0.13) | 0.071 | 0.69* | -0.12 (0.13) | 0.366 | 0.34 |
| Generation length | -0.06 (0.15) | 0.698 | 0.28 | 0.15 (0.15) | 0.337 | 0.38 | 0.04 (0.14) | 0.748 | 0.27 |
| Forest dependency | 0.31 (0.13) | 0.014 | 0.89* | -0.01 (0.14) | 0.930 | 0.27 | 0.15 (0.13) | 0.275 | 0.39 |
| Migratory status | -0.13 (0.15) | 0.378 | 0.35 | 0.04 (0.16) | 0.829 | 0.28 | 0.01 (0.16) | 0.951 | 0.28 |
| Latitude | -0.13 (0.18) | 0.444 | 0.34 | 0.14 (0.16) | 0.375 | 0.36 | 0.15 (0.16) | 0.348 | 0.37 |
| Island dwelling | 0.08 (0.12) | 0.492 | 0.31 | 0.06 (0.13) | 0.611 | 0.29 | 0.05 (0.13) | 0.685 | 0.28 |
| Within-pair range overlap | 0.09 (0.12) | 0.446 | 0.32 | 0.02 (0.13) | 0.877 | 0.26 | 0.09 (0.13) | 0.494 | 0.31 |
| Confamilial sympatry | -0.34 (0.13) | 0.011 | 0.92* | -0.19 (0.13) | 0.172 | 0.49* | 0.04 (0.14) | 0.779 | 0.27 |
| Body mass disparity | 0.21 (0.12) | 0.085 | 0.61* | -0.08 (0.13) | 0.569 | 0.29 | -0.16 (0.14) | 0.245 | 0.42* |
| Bill disparity | 0.05 (0.13) | 0.725 | 0.28 | 0.14 (0.13) | 0.309 | 0.38 | 0.30 (0.14) | 0.027 | 0.82* |
| Climatic disparity | 0.00 (0.12) | 0.992 | 0.26 | 0.12 (0.12) | 0.329 | 0.35 | 0.06 (0.13) | 0.638 | 0.28 |
| Oscine/suboscine | -0.03 (0.12) | 0.778 | 0.26 | -0.09 (0.13) | 0.477 | 0.31 | 0.12 (0.13) | 0.363 | 0.34 |

Table S7. Standardised coefficient estimates from model averaged multipredictor GLMs predicting variation in with in-pair disparity in three independent axes of song variation among passerine species pairs ( $\mathrm{n}=246$ ). SE, standard error; RI, relative importance. * Denotes predictor variables included in the $\mathrm{AIC}_{\mathrm{c}}$ top model.

Full dataset

|  | Pitch disparity (PC1) |  | Length disparity (PC2) |  | Pace disparity (PC3) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate (SE) | $P$ | Estimate (SE) | $P$ | Estimate (SE) | $P$ |
| Plumage dichromatism | -0.34 (0.11) | 0.002 | -0.30 (0.12) | 0.010 | - | - |
| Pair age | 0.31 (0.12) | 0.007 | 0.34 (0.12) | 0.004 | 0.21 (0.12) | 0.095 |
| Body mass | 0.33 (0.11) | 0.002 | 0.27 (0.12) | 0.027 | - | - |
| Generation length | - | - | - | - | - | - |
| Forest dependency | 0.34 (0.12) | 0.005 | - | - | - | - |
| Migratory status | - | - | - | - | - | - |
| Latitude | - | - | - | - | - | - |
| Island dwelling | - | - | - | - | - | - |
| Within-pair range overlap | - | - | - | - | - | - |
| Confamilial sympatry | -0.31 (0.12) | 0.010 | -0.21 (0.12) | 0.075 | - | - |
| Body mass disparity | 0.21 (0.12) | 0.070 | - | - | -0.20 (0.13) | 0.131 |
| Bill disparity | - | - | - | - | 0.30 (0.13) | 0.024 |
| Climatic disparity | - | - | - | - | - | - |
| Oscine/suboscine | - | - | - | - | - | - |

Table S8. Standardised coefficient estimates from $\mathrm{AlC}_{\mathrm{c}}$ top models predicting variation in within-pair disparity in song traits among passerine species pairs ( $n=246$ ). SE, standard error. Pseudo- $R^{2}$ for each model: $\mathrm{PC} 1=0.22 ; \mathrm{PC} 2=0.12 ; \mathrm{PC} 3=0.06$.

Alternative dataset

|  | Model averaged |  |  | $\mathrm{AlC}_{\mathrm{c}}$ top model |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate (SE) | $P$ | RI | Estimate (SE) | $P$ |
| Plumage dichromatism | -0.16 (0.06) | 0.010 | 0.94 | -0.16 (0.06) | 0.001 |
| Pair age | 0.21 (0.06) | 0.001 | 0.99 | 0.20 (0.06) | 0.001 |
| Body mass | 0.10 (0.07) | 0.137 | 0.58 | 0.09 (0.06) | 0.158 |
| Generation length | 0.03 (0.08) | 0.722 | 0.29 | - | - |
| Forest dependency | 0.17 (0.07) | 0.017 | 0.91 | 0.17 (0.06) | 0.009 |
| Migratory status | 0.04 (0.08) | 0.611 | 0.30 | - | - |
| Latitude | 0.06 (0.09) | 0.497 | 0.34 | - | - |
| Island dwelling | 0.01 (0.07) | 0.849 | 0.26 | - | - |
| Within-pair range overlap | 0.06 (0.07) | 0.388 | 0.35 | - | - |
| Confamilial sympatry | -0.15 (0.07) | 0.040 | 0.80 | -0.16 (0.07) | 0.017 |
| Body mass disparity | 0.01 (0.07) | 0.872 | 0.27 | - | - |
| Bill disparity | 0.10 (0.07) | 0.124 | 0.58 | 0.09 (0.06) | 0.134 |
| Climatic disparity | 0.03 (0.07) | 0.647 | 0.28 | - | - |
| Oscine/suboscine | 0.07 (0.07) | 0.298 | 0.40 | - | - |

Table S9. Standardised coefficient estimates from model averaged and AIC $_{c}$ top model multipredictor GLMs predicting variation in within-pair song disparity (PC1-3) using an alternative dataset of passerine species pairs ( $n=246$ ). SE, standard error; RI, relative importance. Pseudo- $R^{2}$ for $\mathrm{AIC}_{\mathrm{c}}$ top model $=0.16$.

Alternative dataset

|  |  |  |  | Alterna | e data |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pitch disparity (PC1) |  |  | Length disparity (PC2) |  |  | Pace disparity (PC3) |  |  |
|  | Estimate (SE) | $P$ | RI | Estimate (SE) | $P$ | RI | Estimate (SE) | $P$ | Rl |
| Plumage dichromatism | -0.15 (0.05) | 0.004 | 0.99* | -0.14 (0.06) | 0.022 | 0.92* | -0.06 (0.06) | 0.288 | 0.41 |
| Pair age | 0.16 (0.05) | 0.004 | 0.99* | 0.14 (0.06) | 0.018 | 0.93* | 0.10 (0.06) | 0.085 | 0.71* |
| Body mass | 0.15 (0.06) | 0.009 | 0.97* | 0.13 (0.07) | 0.056 | 0.82* | -0.06 (0.06) | 0.309 | 0.40 |
| Generation length | -0.01 (0.07) | 0.889 | 0.26 | 0.05 (0.08) | 0.497 | 0.34 | 0.01 (0.07) | 0.888 | 0.27 |
| Forest dependency | 0.16 (0.06) | 0.005 | 0.98* | 0.08 (0.07) | 0.237 | 0.47* | 0.07 (0.06) | 0.264 | 0.44 |
| Migratory status | 0.00 (0.07) | 0.970 | 0.26 | 0.04 (0.08) | 0.599 | 0.32 | 0.01 (0.08) | 0.928 | 0.29 |
| Latitude | -0.01 (0.07) | 0.885 | 0.26 | 0.09 (0.08) | 0.262 | 0.47* | 0.09 (0.07) | 0.210 | 0.51* |
| Island dwelling | 0.04 (0.05) | 0.502 | 0.31 | -0.01 (0.06) | 0.819 | 0.26 | 0.03 (0.06) | 0.579 | 0.30 |
| Within-pair range overlap | 0.02 (0.05) | 0.702 | 0.27 | 0.03 (0.07) | 0.615 | 0.30 | 0.03 (0.06) | 0.597 | 0.29 |
| Confamilial sympatry | -0.17 (0.06) | 0.003 | 0.99* | -0.08 (0.07) | 0.227 | 0.49 | -0.01 (0.07) | 0.911 | 0.27 |
| Body mass disparity | 0.10 (0.06) | 0.078 | 0.69* | -0.04 (0.07) | 0.595 | 0.30 | -0.06 (0.06) | 0.311 | 0.41 |
| Bill disparity | 0.01 (0.06) | 0.810 | 0.27 | 0.07 (0.07) | 0.319 | 0.41 | 0.14 (0.06) | 0.026 | 0.90* |
| Climatic disparity | 0.00 (0.05) | 0.981 | 0.25 | 0.07 (0.06) | 0.291 | 0.42 | 0.02 (0.06) | 0.756 | 0.27 |
| Oscine/suboscine | 0.02 (0.06) | 0.763 | 0.26 | 0.01 (0.06) | 0.869 | 0.26 | 0.09 (0.06) | 0.124 | 0.61* |

Table S10. Standardised coefficient estimates from model averaged multipredictor GLMs predicting variation in within-pair disparity in song traits using an alternative dataset of passerine species pairs ( $\mathrm{n}=246$ ). SE, standard error; RI, relative importance. * Denotes predictor variables included in the $\mathrm{AIC}_{\mathrm{c}}$ top model.

|  | Alternative dataset |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pitch disparity (PC1) |  | Length disparity (PC2) |  | Pace disparity (PC3) |  |
|  | Estimate (SE) | $P$ | Estimate (SE) | $P$ | Estimate (SE) | $P$ |
| Plumage dichromatism | -0.15 (0.05) | 0.003 | -0.14 (0.06) | 0.018 | - | - |
| Pair age | 0.15 (0.05) | 0.004 | 0.14 (0.06) | 0.018 | 0.09 (0.06) | 0.116 |
| Body mass | 0.14 (0.05) | 0.007 | 0.14 (0.06) | 0.021 | - | - |
| Generation length | - | - | - | - | - | - |
| Forest dependency | 0.16 (0.05) | 0.003 | 0.09 (0.07) | 0.175 | - | - |
| Migratory status | - | - | - | - | - | - |
| Latitude | - | - | -0.13 (0.07) | 0.055 | 0.08 (0.06) | 0.145 |
| Island dwelling | - | - | - | - | - | - |
| Within-pair range overlap | - | - | - | - | - | - |
| Confamilial sympatry | -0.17 (0.05) | 0.002 | - | - | - | - |
| Body mass disparity | 0.10 (0.05) | 0.066 | - | - | - | - |
| Bill disparity | - | - | - | - | 0.12 (0.06) | 0.036 |
| Climatic disparity | - | - | - | - | - | - |
| Oscine/suboscine | - | - | - | - | 0.09 (0.06) | 0.098 |

Table S11. Standardised coefficient estimates from AIC $_{c}$ top models predicting variation in within-pair disparity in song traits using an alternative dataset of passerine species pairs ( $\mathrm{n}=$ 246). SE, standard error. Pseudo- $R^{2}$ for each model: $P C 1=0.23 ; ~ P C 2=0.12 ; ~ P C 3=0.08$.


Table S12. Pearson correlation coefficients ( $r^{2}$ ) for all pairwise combinations of (standardised) predictor variables considered in our analysis

